

NOTES ON CONTINUOUS RANDOM WALKS
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The stuff below is standard but not so easy to extract from the literature.

1. CONTINUOUS RANDOM WALKS

We discuss a continuous 1-dimensional random walk with drift μ and variance σ^2 per time unit. We assume there is some boundary C in the $x-t$ -plane (t is the time coordinate and x is a spatial coordinate) such that if the random walk touches C there is a payoff of $\psi(x, t)$.

Let $P(x, t)$ be the expectation value of the eventual payoff for a particle at (x, t) . Then the value of $P(x, t)$ is governed by the following diffusion equation

$$(1.1) \quad \frac{\partial P}{\partial t} = -\mu \frac{\partial P}{\partial x} - \frac{\sigma^2}{2} \frac{\partial^2 P}{\partial x^2}$$

2. STOPPING TIME

The stopping time S is defined as the first time that the random walk touches the boundary of an interval of length A . Below we compute the probability density of S .

To do this we first compute the probability that S is larger than a given constant T . I.e. $P(S \geq T)$. To do this have to solve (1.1) with the boundary conditions $P(0, t) = P(A, t) = 0$ for $t < T$ and $P(x, T) = 1$. We will look at the more general problem where $P(x, T) = \psi(x)$.

Using separation of variables $P(x, t) = X(x)T(t)$ we have to solve

$$\begin{aligned} \frac{\partial T}{\partial t} &= \lambda T \\ \frac{\sigma^2}{2} \frac{\partial^2 X}{\partial x^2} + \mu \frac{\partial X}{\partial x} + \lambda X &= 0 \end{aligned}$$

with $X(0) = 0$, $X(A) = 0$.

The roots of the characteristic equation of the last equation are

$$\chi_{1,2} = \frac{-\mu \pm \sqrt{\mu^2 - 2\sigma^2\lambda}}{\sigma^2}$$

It follows that we should have

$$\Delta = 2\sigma^2\lambda - \mu^2 > 0$$

and the general solutions are

$$e^{-x\mu/\sigma^2} \sin x\sqrt{\Delta}/\sigma^2$$

as the cosine solutions are excluded by our boundary condition $X(0) = 0$. The boundary condition $X(A) = 0$ leads to

$$A\sqrt{\Delta}/\sigma^2 = n\pi$$

which leads to the following values for λ :

$$\lambda_n = \frac{n^2\pi^2\sigma^2}{2A^2} + \frac{\mu^2}{2\sigma^2}$$

Hence our general solution is now

$$e^{\lambda_n t - x\mu/\sigma^2} \sin n\pi x/A$$

We look for a_n such that

$$\sum_n a_n e^{\lambda_n T - x\mu/\sigma^2} \sin n\pi x/A = \psi(x)$$

or equivalently b_n such that

$$\sum_n b_n \sin n\pi x/A = e^{\gamma x} \psi(x)$$

where $\gamma = \mu/\sigma^2$ and $b_n = a_n e^{\lambda_n T}$. Standard Fourier analysis yields

$$b_n = \frac{2}{A} \int_0^A e^{\gamma y} \sin \frac{n\pi y}{A} \psi(y) dy$$

So the solution is

$$(2.1) \quad P(x, t) = \sum_n e^{-\lambda_n t} e^{\lambda_n T - \gamma x} \sin \frac{n\pi x}{A} \left[\frac{2}{A} \int_0^A e^{\gamma y} \sin \frac{n\pi y}{A} \psi(y) dy \right]$$

Our original problem corresponds to

$$\psi(x) = \begin{cases} 1 & 0 \leq x \leq A \\ 0 & \text{otherwise} \end{cases}$$

Thus the probability that the stopping time S is $\geq T$ is

$$P(S \geq T) = \sum_n e^{-\lambda_n T} e^{-\gamma x} \sin \frac{n\pi x}{A} \left[\frac{2}{A} \int_0^A e^{\gamma y} \sin \frac{n\pi y}{A} dy \right]$$

For the sequel it will be convenient to define

$$\begin{aligned} H_n(\gamma, x) &= \frac{2}{A} \int e^{\gamma x} \sin \frac{n\pi x}{A} dx \\ &= \frac{2A\gamma e^{\gamma x} \sin \pi n x/A - 2\pi n e^{\gamma x} \cos \pi n x/A}{A^2\gamma^2 + \pi^2 n^2} \end{aligned}$$

In particular we have the following special values

$$\begin{aligned} H_n(\gamma, 0) &= -\frac{2\pi n}{A^2\gamma^2 + \pi^2 n^2} \\ H_n(\gamma, A) &= -(-1)^n \frac{2\pi n e^{\gamma A}}{A^2\gamma^2 + \pi^2 n^2} \end{aligned}$$

so that we find

$$P(S \geq T) = \sum_n e^{-\lambda_n T - \gamma x} (H_n(\gamma, A) - H_n(\gamma, 0)) \sin \frac{n\pi x}{A}$$

From this we compute the expected stopping time. The probability distribution of S is

$$p(S = T) = \sum_n \lambda_n e^{-\lambda_n T - \gamma x} (H_n(\gamma, A) - H_n(\gamma, 0)) \sin \frac{n\pi x}{A}$$

so that

$$E(S) = \sum_n \frac{1}{\lambda_n} e^{-x\gamma} (H_n(\gamma, A) - H_n(\gamma, 0)) \sin \frac{n\pi x}{A}$$

Assume now that the experiment is truncated at a certain fixed time T . Then the modified stopping time is

$$S' = \begin{cases} S & \text{if } S \leq T \\ T & \text{if } S \geq T \end{cases}$$

Taking into account that

$$\int_0^T \lambda t e^{-\lambda t} dt = -\frac{(\lambda t + 1)e^{-\lambda t}}{\lambda} \Big|_0^T = \frac{1}{\lambda} (1 - (T\lambda + 1)e^{-T\lambda})$$

we find

$$E(S') = P(S \geq T)T + \sum_n \frac{1}{\lambda_n} (1 - (T\lambda_n + 1)e^{-\lambda_n T}) e^{-x\gamma} (H_n(\gamma, A) - H_n(\gamma, 0)) \sin \frac{n\pi x}{A}$$

3. THE PROBABILITY THAT THE PARTICLE HITS $t = T$, $x \leq y$ BEFORE HITTING THE UPPER OR LOWER BOUNDARY

In this case we must use (2.1) with

$$\psi(x) = \begin{cases} 1 & x \leq y \\ 0 & x > y \end{cases}$$

Thus we find

$$P(x, t) = \sum_n e^{-\lambda_n t} e^{\lambda_n x} (H(\gamma, y) - H(\gamma, 0)) \sin \frac{n\pi x}{A}$$

4. THE PROBABILITY THAT THE PARTICLE TOUCHES THE UPPER BOUNDARY BEFORE TIME T

We must solve the boundary value problem for P with $P(0, t) = P(x, T) = 0$, $P(A, t) = 1$.

We first look for a $Q(x, t)$ solution independent of t which satisfies $Q(0, t) = 0$, $Q(A, t) = 1$. Such a solution is of the form

$$C + D e^{-2\gamma x}$$

for suitable constants C, D . The boundary conditions give the following constraints

$$\begin{aligned} C + D &= 0 \\ C + D e^{-2\gamma A} &= 1 \end{aligned}$$

which leads to

$$\begin{aligned} C &= -\frac{1}{e^{-2\gamma A} - 1} \\ D &= \frac{1}{e^{-2\gamma A} - 1} \end{aligned}$$

Hence the solution has the form

$$Q(x, t) = \frac{e^{-2\gamma x} - 1}{e^{-2\gamma A} - 1}$$

Consider now

$$P'(x, t) = P(x, t) - Q(x, t)$$

Then $P'(0, t) = P'(A, t) = 0$ for $t \leq A$ and

$$P'(x, T) = -\frac{e^{-2\gamma x} - 1}{e^{-2\gamma A} - 1}$$

so that using (2.1)

$$P'(x, t) = -\frac{1}{e^{-2\gamma A} - 1} \sum_n e^{-\lambda_n T} e^{\lambda_n t - \gamma x} \sin \frac{n\pi x}{A} \left[\frac{2}{A} \int_0^A e^{\gamma y} \sin \frac{n\pi y}{A} (e^{-2\gamma y} - 1) dy \right]$$

So the ultimate solution is

$$P(x, t) = \frac{e^{-2\gamma x} - 1}{e^{-2\gamma A} - 1} - \frac{1}{e^{-2\gamma A} - 1} \sum_n e^{-\lambda_n T} e^{\lambda_n t - \gamma x} \sin \frac{n\pi x}{A} \left[\frac{2}{A} \int_0^A e^{\gamma y} \sin \frac{n\pi y}{A} (e^{-2\gamma y} - 1) dy \right]$$

which can also be written as

$$P(x, t) = \frac{e^{-2\gamma x} - 1}{e^{-2\gamma A} - 1} - \frac{1}{e^{-2\gamma A} - 1} \sum_n e^{-\lambda_n T} e^{\lambda_n t - \gamma x} \sin \frac{n\pi x}{A} (H_n(-\gamma, A) - H_n(\gamma, A))$$

To get the probability that the particle touches the lower boundary first we have to make the substitutions $x \mapsto A - x$, $\gamma \mapsto -\gamma$.